

# Fast methods for Bayesian inverse problems with uncertain PDE forward models

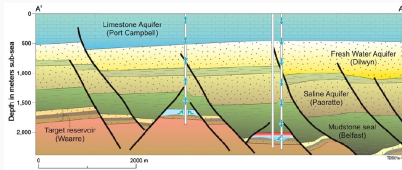
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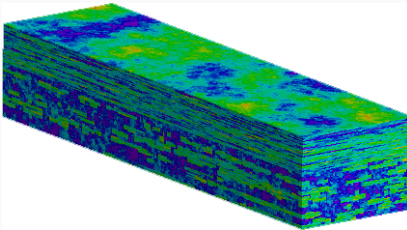
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# Example of inverse problem with uncertain forward model

Infer **fault transmissibility** in a subsurface flow model:



Berkeley Lab, [newscenter.lbl.gov](http://newscenter.lbl.gov)



SPE 10, [spe.org/web/csp](http://spe.org/web/csp)

- PDE model: **multi-phase porous medium flow** (expensive to solve)
- Observations: **fluid pressure and well production** (typically sparse and noisy)
- Presence of random parameter field: **background permeability** (reconstructed from previous studies)

## Derivation of the algorithm

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# Inverse problem governed by random PDE forward problem

Infer the **inversion parameter**  $m \in \mathcal{M}$ , given noisy observations  $d \in V$  of the state variable  $u \in \mathcal{U}$ , governed by the random PDE

$$r(u, m, k) = 0$$

where

- **Random parameter**  $k \in \mathcal{K}$  is normally distributed,  
 $k \sim \mathcal{N}(\bar{k}, \Gamma_k)$
- Observations are polluted by Gaussian additive noise,  
 $d = \mathcal{B}u(k, m) + \beta, \quad \beta \sim \mathcal{N}(0, \Gamma_\beta)$
- Prior information about  $m$  is given by the measure  $\mu_{\text{pr}}$

# Bayesian solution of inverse problem

By (infinite dimensional) Bayes' Theorem,

$$\frac{d\mu_{\text{post}}}{d\mu_{\text{pr}}} = \frac{1}{\mathcal{Z}} \pi_{\text{like}}(d|m)$$

Marginalize out the random parameter  $k$ :

$$\begin{aligned} \pi_{\text{like}}(d|m) &\propto \int_{\mathcal{K}} \pi(d|m, k) d\mu_k \\ &= \int_{\mathcal{K}} \exp \left[ -\frac{1}{2} (d - \mathcal{B}u(m, k))^T \Gamma_{\beta}^{-1} (d - \mathcal{B}u(m, k)) \right] d\mu_k \end{aligned}$$

For sufficiently complex forward models, this integral is **prohibitive** to approximate via Monte Carlo sampling (and it has to be done for each sample of  $m$ )

## Approximation of the likelihood

Consider the **first-order approximation** of  $u$  wrt  $k$ :

$$u^L(m, k) = u(m, \bar{k}) + D_k u(m, \bar{k})[k - \bar{k}]$$

where  $D_k$  is the Fréchet derivative of  $u$  wrt  $k$

The likelihood based on  $u^L$  becomes:

$$\pi_{\text{like}}^L \propto \exp\left(\frac{1}{2}\|d - \mathcal{B}u(m, \bar{k})\|_{\Gamma_\nu^{-1}}^2 - \frac{1}{2}\log \det(\Gamma_\nu)\right)$$

where  $\Gamma_\nu := \Gamma_\beta + (\mathcal{B}D_k u)\Gamma_k(\mathcal{B}D_k u)^T$

The term  $(\mathcal{B}D_k u)\Gamma_k(\mathcal{B}D_k u)^T$  is the contribution of model uncertainty to the noise covariance, and results from pushing forward the covariance of  $k$ ,  $\Gamma_k$ , through the Jacobian of the parameter-to-observable map,  $\mathcal{B}D_k u$ .

# Low rank approximation of sensitivity matrix

- As  $m$  changes across optimization iterations/MCMC sampling steps, we need to evaluate (and later compute derivatives wrt  $m$  of) the term

$$\Gamma_\beta + \underbrace{J(m, \bar{k})}_{(BD_k u)} \Gamma_k \underbrace{J(m, \bar{k})^T}_{(BD_k u)^T}$$

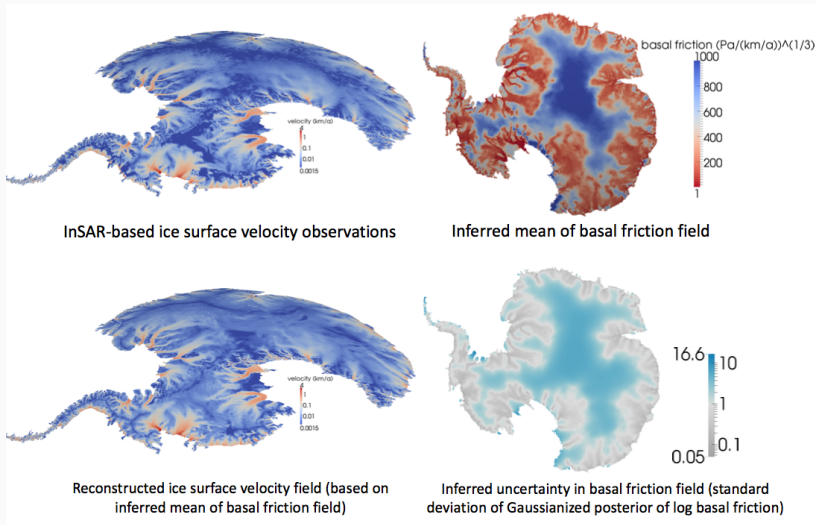
- $J \Gamma_k J^T$  resembles the Gauss-Newton Hessian (of the log-likelihood wrt  $k$ ), but turned inside-out (i.e.  $J J^T$  operating on data space instead of  $J^T J$  operating on parameter space)
- Explicit construction of  $J$  requires as many (linearized) forward PDE solves as  $\min(\text{data dimension}, k \text{ dimension})$ , which is prohibitive
- For many problems, singular values of  $BD_k u$  decay rapidly and a (randomized) low rank approximation can be made:

$$J \Gamma_k J^T \approx V_r \Lambda_r V_r^T$$

where the columns of  $V_r$  contain the  $r$  eigenvectors of  $J \Gamma_k J^T$  and the diagonal elements of  $\Lambda_r$  contains its eigenvalues

- Use **randomized SVD** which requires  $O(r)$  matrix-vector products with random vectors  $\xi_i$ , amounting to  $O(r)$  linearized adjoint solves  $\eta_i = J^T \xi_i$  and linearized forward solves  $J \Gamma_k \eta_i$

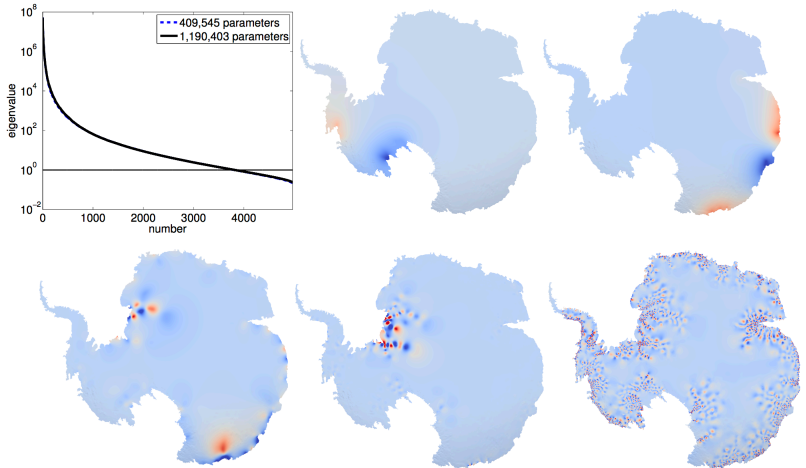
# Example: Antarctic ice sheet flow inversion



T. Isaac, N. Petra, G. Stadler, O. Ghattas, JCP 2015



# Spectrum of Hessian for Antarctic ice flow inverse problem



$O(i^{-3})$  eigenvalue decay of prior-preconditioned data misfit Hessian (5000 out of 1.19 million) and eigenvectors 1, 7, 100, 500, 4000

## Eigenproblem-constrained optimization for MAP point

With the low-rank approximation, the **maximum a posteriori** estimate,  $m_{\text{MAP}}$ , is found by minimizing the negative log of the posterior distribution:

$$m_{\text{MAP}} = \arg \min_{m \in \mathcal{M}} \left( \frac{1}{2} \|\mathcal{B}u(m, \bar{k}) - d\|_{(\Gamma_\beta + V_r \Lambda_r V_r^T)^{-1}}^2 + \frac{1}{2} \|\Gamma_{\text{pr}}^{-\frac{1}{2}}(m - \bar{m})\|_{L^2(\Omega)}^2 + \frac{1}{2} \log \det(\Gamma_\beta + V_r \Lambda_r V_r^T) \right)$$

where the state solution  $u$  depends on the parameters  $m, \bar{k}$  through

$$(r(u, m, \bar{k}), \tilde{u}) = 0 \quad \forall \tilde{u} \in \mathcal{U}$$

and the eigenvalues  $\lambda_i$  & eigenvectors  $v_i$  depend on  $u, m, \bar{k}$  through

$$(J(\bar{k}, m, u) \Gamma_k J(\bar{k}, m, u)^T v_i, \tilde{v}_i) = \lambda_i (v_i, \tilde{v}_i) \quad \forall \tilde{v}_i \in V, i = 1, \dots, r$$

- Solving the eigenvalue problem via randomized SVD amounts to solving  $O(r)$  pairs of linearized adjoint/forward PDEs
- Computing the gradient requires differentiating through the eigenvalue problem to obtain eigenvalue and eigenvector sensitivities, necessitating an solution of  $O(r)$  adjoint eigenvalue problems.

## Current simpler implementation based on MC estimator

- Have constructed fast algorithms for Hessian-constrained optimization in other contexts (e.g., P. Chen, U. Villa, and O. Ghattas, JCP 2019)
- But somewhat complicated implementation, so initially consider a Monte Carlo estimator of  $J\Gamma_k J^T$
- Let  $e$  be a standard normal random variable; then

$$\begin{aligned}(\mathcal{B}D_k u(m, \bar{k}))\Gamma_k(\mathcal{B}D_k u(m, \bar{k}))^T &= (\mathcal{B}D_k u(m, \bar{k}))\Gamma_k^{1/2}\mathbb{E}[ee^T]\Gamma_k^{1/2}(\mathcal{B}D_k u(m, \bar{k}))^T \\ &= \mathbb{E}[(\mathcal{B}D_k u(m, \bar{k}))\Gamma_k^{1/2}e(\mathcal{B}D_k u(m, \bar{k}))\Gamma_k^{1/2}e^T]\end{aligned}$$

- Consider a Monte Carlo estimator with  $\xi_i \sim \mathcal{N}(0, \Gamma_k)$ ,  $i = 1, \dots, n_s$ ,

$$(\mathcal{B}D_k u(m, \bar{k}))\Gamma_k(\mathcal{B}D_k u(m, \bar{k}))^T \approx \frac{1}{n_s} \sum_{i=1}^{n_s} (\mathcal{B}D_k u(m, \bar{k})[\xi_i]) (\mathcal{B}D_k u(m, \bar{k})[\xi_i])^T$$

## MC estimator-based MAP point computation

The **maximum a posteriori** estimate,  $m_{\text{MAP}}$ , is found by minimizing the negative log of the posterior distribution:

$$m_{\text{MAP}} = \arg \min_{m \in \mathcal{M}} \left( \frac{1}{2} \|\mathcal{B}u(m, \bar{k}) - d\|_{(\Gamma_\beta + \frac{1}{n_s} \sum_{i=1}^{n_s} \mathcal{B}U_i \mathcal{B}U_i^T)^{-1}}^2 + \frac{1}{2} \|\Gamma_{\text{pr}}^{-\frac{1}{2}}(m - \bar{m})\|_{L^2(\Omega)}^2 + \frac{1}{2} \log \det(\Gamma_\beta + \frac{1}{n_s} \sum_{i=1}^{n_s} \mathcal{B}U_i \mathcal{B}U_i^T) \right)$$

where the state solution  $u$  depends on the parameters  $m, \bar{k}$  through

$$(r(u, m, \bar{k}), \tilde{u}) = 0 \quad \forall \tilde{u} \in \mathcal{U}$$

and the sensitivity solutions,  $U_i$ , depend on  $u, m, \bar{k}$  through

$$((\partial_u r)(\bar{k}, m, u)[U_i], \tilde{U}_i) = -((\partial_k r)(\bar{k}, m, u)[\xi_i], \tilde{U}_i) \quad \forall \tilde{U}_i \in \mathcal{U}, i = 1, \dots, n_s$$

This is an  $n_s + 1$  PDE constrained optimization problem, but it can be solved efficiently using gradient based optimization, since  $n_s$  is typically small and independent of the parameter dimension. This requires an efficient computation of the gradient...

# Gradient computation for the MAP optimization problem

The (weak form of the) **gradient** is given by

$$\begin{aligned} (\mathcal{G}(m), \tilde{m}) &= (m - \bar{m}, \Gamma_{pr}^{-1} \tilde{m}) + ((\partial_m r)(\bar{k}, m, u)[\tilde{m}], v) \\ &+ \sum_{i=1}^{n_s} [((\partial_{u_m} r)(\bar{k}, m, u)[U_i, \tilde{m}], V_i) + ((\partial_{k_m} r)(\bar{k}, m, u)[\xi_i, \tilde{m}], V_i)], \forall \tilde{m} \in \mathcal{M} \end{aligned}$$

where  $u$  and  $\{U_i\}_{i=1}^{n_s}$  solve the “forward” problems

$$\begin{aligned} (r(u, m, \bar{k}), \tilde{u}) &= 0 \quad \forall \tilde{u} \in V \\ ((\partial_u r)(\bar{k}, m, u)[U_i, \tilde{U}_i] &= -((\partial_k r)(\bar{k}, m, u)[\xi_i, \tilde{U}_i]) \quad \forall \tilde{U}_i \in V, i = 1, \dots, n_s \end{aligned}$$

and  $v$  and  $\{V_i\}_{i=1}^{n_s}$  solve the “adjoint” problems

$$\begin{aligned} ((\partial_u r)(u, m, \bar{k})[\tilde{v}], v) &= -(\mathcal{B}u - d, \Gamma_\nu^{-1} \tilde{v}) \\ &- \sum_{i=1}^{n_s} [((\partial_{uu} r)(\bar{k}, m, u)[U_i, \tilde{v}], V_i) + ((\partial_{ku} r)(\bar{k}, m, u)[\xi_i, \tilde{v}], V_i)], \forall \tilde{v} \in V \\ ((\partial_{uU_i} r)(\bar{k}, m, u)[U_i, \tilde{V}_i], V_i) &+ ((\partial_{kU_i} r)(\bar{k}, m, u)[\xi_i, \tilde{V}_i], V_i) = -\frac{1}{n_s} (\mathcal{B}^T \Gamma_\nu^{-1} \mathcal{B}U_i, \tilde{V}_i) \\ &+ \frac{1}{n_s} (\mathcal{B}^T \Gamma_\nu^{-1} (\mathcal{B}u - d)(\mathcal{B}u - d)^T \Gamma_\nu^{-1} \mathcal{B}U_i, \tilde{V}_i), \forall \tilde{V}_i \in V, i = 1, \dots, n_s \end{aligned} \quad 11$$

## Numerical examples

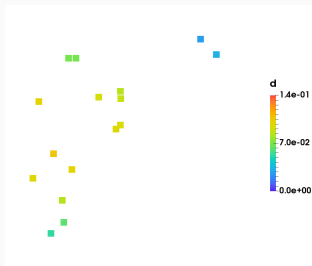
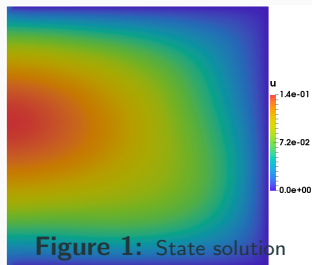
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## Test case 1: Poisson problem

The forward problem: find  $u \in V$  s.t.

$$\begin{aligned} -\nabla \cdot (e^m \nabla) u &= k, & \text{in } \Omega, \\ u &= 0, & \text{on } \Gamma_D, \\ e^m \nabla u \cdot n &= 0, & \text{on } \Gamma_N. \end{aligned}$$

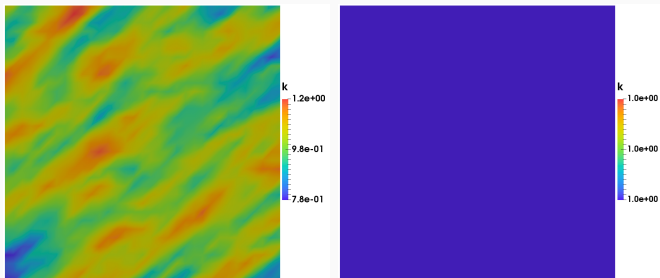
- The domain is  $[0, 1] \times [0, 1]$ ;
- Dirichlet BC on  $\Gamma_D = \{0\} \times [0, 1]$ ;
- 4225 dofs in state space,  
1089 dofs in  $(m, k)$  parameter space
- 16 observations with random locations;
- $n_s = 30$  samples;
- 1% Gaussian noise.



## Test case 1: Random parameter

We assume that  $\mu_k \sim \mathcal{N}(\bar{k}, \Gamma_k)$ , where the covariance is given by  $\Gamma_k = \mathcal{A}^{-2}$   
s.t.  $\mathcal{A}k = \tilde{k}$  satisfies

$$\gamma_k(\Theta_k \nabla k, \nabla v) + \delta_k(k, v) = (\tilde{k}, v) \quad \forall v \in H^1(\Omega)$$



**Figure 3:** Left: sample from  $\mu_k$  (used to synthesize data); right: mean of  $\mu_k$  (used for deterministic inversion)

The prior is given by a distribution with similar structure



## Test case 1: Some comments

- Random parameter-to-state map is linear, so approximation is exact
- Small number of observations, not evenly distributed over domain
- Relatively high variance of random parameter distribution, dominates observational noise

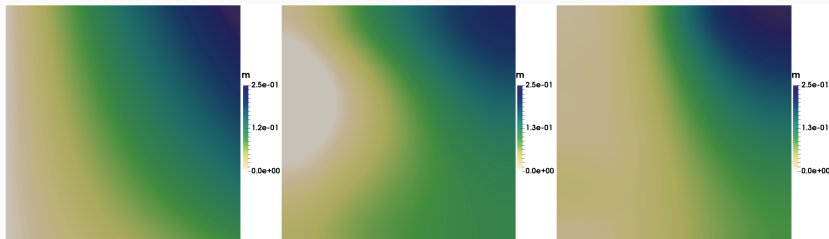
Algorithm has been implemented using [FEniCS](https://fenicsproject.org/)<sup>1</sup> and [hippylib](https://hippylib.github.io/)<sup>2</sup> open source libraries

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<sup>1</sup><https://fenicsproject.org/>

<sup>2</sup><https://hippylib.github.io/>

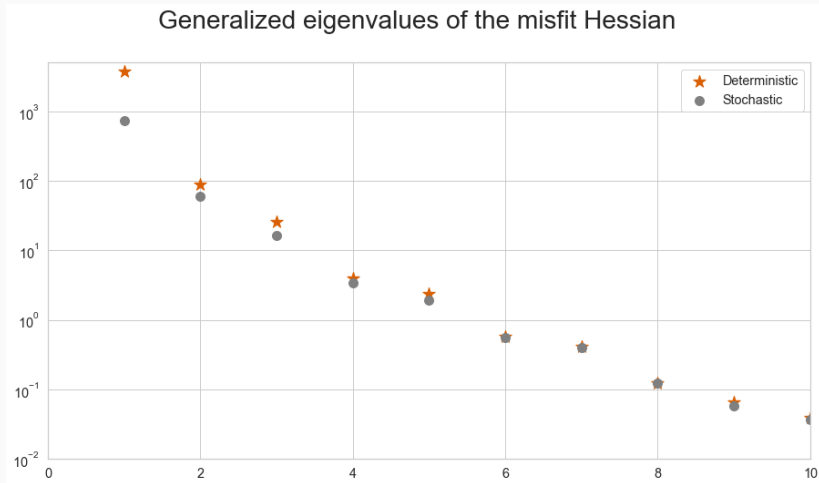
## Test case 1: MAP point



**Figure 4:** Left: true  $m$ ; middle: MAP point  $m^{\text{det}}$ ; right: MAP point  $m^{\text{stoch}}$

We compare the true field  $m$ , the MAP estimate obtained using the deterministic algorithm  $m^{\text{det}}$ , and the MAP point obtained by the proposed algorithm  $m^{\text{stoch}}$ .

# Test case 1: Eigenvalue of data misfit Hessian (wrt $m$ )



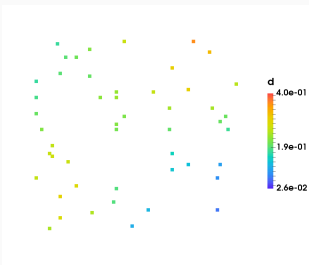
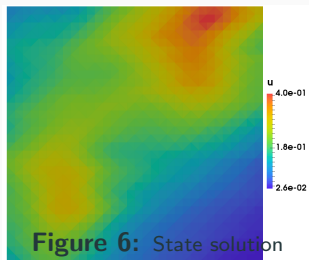
**Figure 5:** Generalized eigenvalues of prior preconditioned data misfit Hessian (wrt  $m$ )

## Test case 2: Advection–diffusion–reaction problem

The forward problem: find  $u \in V$  s.t.

$$\begin{aligned} -\nabla \cdot (\kappa \nabla u) + \nabla u \cdot k + cu &= m, & \text{in } \Omega, \\ u &= 0, & \text{on } \Gamma_D. \end{aligned}$$

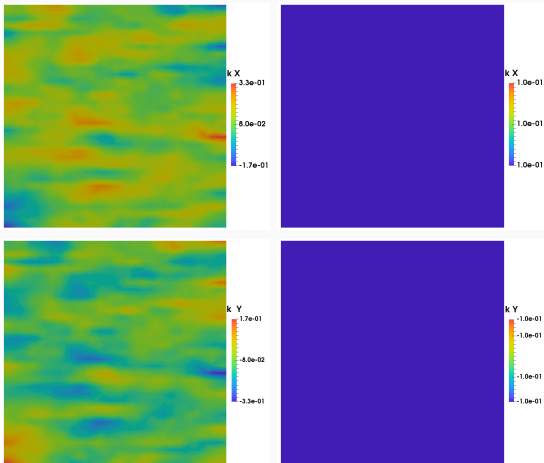
- The domain is  $[0, 1] \times [0, 1]$ ;
- Dirichlet BC on  $\Gamma_D = \{0\} \times [0, 1] \cup [0, 1] \times \{0\}$ ;
- 2048 dofs in state space,  
1089 dofs in parameter space
- 50 observations with random locations;
- $n_s = 30$  samples;
- 1% Gaussian noise.



## Test case 2: Random parameter

We assume that  $\mu_k \sim \mathcal{N}(\bar{k}, \Gamma_k)$ , where covariance given by  $\Gamma_k = \mathcal{A}^{-2}$  s.t.  $\mathcal{A}k = \tilde{k}$  satisfies

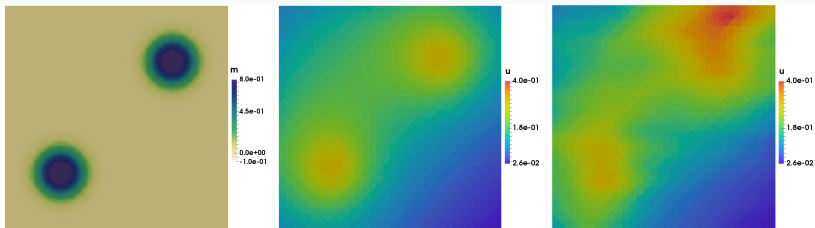
$$\gamma_k(\Theta_k \nabla k, \nabla v) + \delta_k(k, v) = (\tilde{k}, v), \quad \forall v \in H^1(\Omega).$$



**Figure 9:** Left: sample from  $\mu_k$ ; right: mean of  $\mu_k$ .

## Test case 2: State solution

The state solution is sensitive to the random parameter



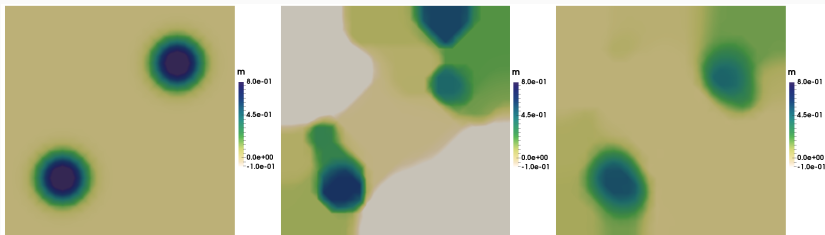
**Figure 10:** Left: true  $m$ ; middle:  $u(m, \bar{k})$ ; right:  $u(m, k)$ .

Diffusion coefficient  $\kappa = 10^{-2}$ , reaction coefficient  $c = 0.4$

## Test case 2: Some comments

- Random parameter-to-state map is not linear, so approximation is not exact
- Moderate number of observations, evenly distributed over the domain
- Large variance in random parameter distribution, dominates noise
- Sharp interfaces in the inversion parameter field (TV prior used)

## Test case 2: MAP point



**Figure 11:** Left: true  $m$ ; middle: MAP point  $m^{\text{det}}$ ; right: MAP point  $m^{\text{stoch}}$ .

We compare the true field  $m$ , the MAP estimate obtained using traditional algorithm  $m^{\text{det}}$ , and the MAP point obtained by the proposed algorithm  $m^{\text{stoch}}$ .



## Conclusions

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# Conclusions

We considered the Bayesian inference of an unknown parameter field  $m$  in the presence of uncertainty in another parameter field  $k$

- We proposed a method based on a linearized approximation of the map from  $k$  to  $u$  that accounts for both the uncertainty and the sensitivity of the state wrt  $k$
- Computation of the negative log posterior and its gradient require a number of PDE solves ( $n_s+1$ ) that is small and independent of the dimensions of the unknown parameter  $m$  and the random parameter  $k$
- Examples illustrate the effect of accounting for model uncertainty in the Bayesian solution

## Ongoing and future work

- Approximation of the Jacobian  $J$  via randomized truncated SVD (greater accuracy for fewer  $n_s$ )
  - Requires addition of eigenvalue problem with forw/adjoint PDEs
- Comparison of our approach to full MCMC + inner MC solution
- Extension of the method to allow for more general distributions  $\mu_k$  of the random parameter
- Use of the approximation as a control variate:

$$\begin{aligned}\pi_{\text{like}}(d|m) &\propto \mathbb{E}_{\mu_k} [\pi(d|m, k)] \\ &= \mathbb{E}_{\mu_k} [\pi^L(d|m, k)] + \mathbb{E}_{\mu_k} [\pi(d|m, k) - \pi^L(d|m, k)] \\ &\approx \exp\left(\frac{1}{2}\|d - \mathcal{B}u(m, \bar{k})\|_{\Gamma_\nu^{-1}}^2 - \frac{1}{2}\log \det(\Gamma_\nu)\right) \\ &\quad + \frac{1}{N} \sum_{i=1}^N \exp\left(\frac{1}{2}\|d - \mathcal{B}u(m, k_i)\|_{\Gamma_\beta^{-1}}^2\right) \\ &\quad - \frac{1}{N} \sum_{i=1}^N \exp\left(\frac{1}{2}\|d - \mathcal{B}u(m, \bar{k}) - \mathcal{B}D_k u(m, \bar{k})[k_i - \bar{k}]\|_{\Gamma_\beta^{-1}}^2\right)\end{aligned}$$